## Using Rasch Model for the Calibration of Test Items in Mathematics, Grade-9

<sup>1</sup> Muhammad Sadiq, <sup>2</sup> Shamim Haider Tirmizi. <sup>3</sup> Mubashra Jamil

<sup>1</sup> Research Scholar Ph.D., Department of Education BZU Multan
 <sup>2</sup> Associate Prof., Department of Education BZU Multan
 <sup>3</sup> Assistant Prof. Department of Education BZU Multan

(Email: m.sadiqap@gmail.com)

The main objective of the study was the development of a semi-standardized test in Mathematics at grade 9 through Rasch modelling. The population of the study was 59,168 students of grade 9 admitted to 718 secondary schools of Bahawalpur Division (Pakistan) during session 2010-11. The schools were categorised gender-wise, locality-wise and ownership-wise to ensure representation of the population in the sample. Cluster sampling was used to identify the sample. For this, 3 schools were randomly opted from each category and all the students of grade 9 in these schools were included in the sample. There were 720 students. However, 642 students participated in the study due to different reasons. Two equivalent test forms, each with 43 items were developed from Mathematics of grade 9. These were ultimately transformed into a 86-items test to be semi-standardized. The content validity of the test was obtained by incorporating the opinion of teachers in Mathematics and researchers. The alternate-forms reliability of the test was 0.92. The two test forms were administered everywhere with uniformity of instructions and arrangements for two consecutive days to retain the external validity of the test. The Rasch analysis showed more accuracy thoroughly in the appraisal of test items than respondents in terms of logit values, Model error estimates and fit-statistics. The major recommendations of the study were the replication of the current study, its execution at other times during the academic session, and finding probability correct for all items and persons to see response patterns more realistically

Key words: Rasch analysis, difficulty and ability measures, precision of estimation, fit-statistics

#### Introduction

Assessment is an integral part of formal education with good intentions to appraise and monitor performance on academic tasks. This activity shapes both curriculum and teaching (Stobart & Gipps, 1997). It is regrettable that mostly, assessment fails to fulfill its intended outcomes and involve students in short cuts. Stobart (2008) is right to say that current assessment techniques promote shallow knowledge, put students to focus on end points and cram curriculum. Usually, ill-defined objectives, poorly-designed tests and superficial expectations make assessment objectionable and cause its abuses. Stobart (2008) sees assessment differently as a value-laden social activity, a procedure to create and plan things to be measured, and a way to determine the contents and ways of learning. The author's view-point with all its essence appears to open new horizons to explore ways to

shift assessment from "having mode" to "being mode." Testing is a way of assessment. It is an enormous enterprise and mostly relies on classroom and standardized testing.

Haladyna and Rodriguez (2013) say that both aspects of testing employ the same theories and technology with somewhat different objectives. In Pakistan, large-scale/standardized testing is not very popular at school and college level. There is need to proceed in this direction for across the classroom comparisons on certain academic tasks.

No doubt, the concept of standardization is commonly associated with psychological testing, yet it is equally used to appraise students on academic tasks in different content areas. No definition of test standardization restricts it to psychological perspective rather opens it for educational use. Slavin (2012, p.448) supports the view point defining standardized tests in educational context as, "tests that are usually commercially prepared for nationwide use and designed to provide accurate and meaningful information on students' performance relative to that of others at their age or grade levels." Kubiszyn and Borich (2010) confirm the claim that standardized achievement tests are developed by test construction experts with the assistance of curriculum developers and classroom teachers to compare a student's achievement with his age and grade cohorts. Generally, standardized group tests cover language skills, mathematics, science and social studies. In the same perspective, Thorndike and Hagan (1977) disclose that such tests may focus on a particular subject or even on a particular course at secondary school level. Mrunalini (2011) characterises that items for standardized tests are expertly written, properly pretested, rigorously analysed and scientifically refined, what be the contents. The writer seems to ignore prominent features of standardized testing such as nationwide scope, uniform administration. norms for interpretation, preparation of item profile and age or grade based comparison. Especially, the word "scientifically refined" creates ambiguity and makes no contextual sense. It needs to know, what Mrunalini means it.

Standardized testing is an integral part of school and college education in many countries. Kubiszyn and Borich (2010) reveal that standardized tests are administered to 140-400 million students worldwide yearly. Woolfolk (2011) exposes that all the 50 states and District of Columbia in America have policies on statewide testing. It would be rare if a student has not experienced several standardized tests during school days. Ormord (2012) agrees that use of standardized achievement tests has expanded drastically to assess students and teachers during the recent years. Adding more, Linn and Gronlund (2005)bring forth that norm-referenced standardized tests dominate testing in education either as a part of broader assessment system or alone, being an efficient and relatively cheap mode

for assessing broader achievement goals. Narrating such goals, Earl (2003) reveals that these tests are profitable for accountability, evaluation, comparison, addressing educational placement, concerns. reviewing educational practices, diagnosing students' problems, highlighting strengths and weaknesses of specific programs, revising curricula and rating teachers in addition to compulsion to do so. Earl seems to overrate standardized tests illogically, when expects such a wide range of outcomes from these tests. It is similar to use same prescription to cure from all diseases. Standardized tests alone appear helpless to actualize many of these intentions. Even, these will be inappropriate for some results. The statement by AERA (2000, p.127) as quoted by Crisp (2007, p.49) clarifies the position as "performance on standardized test should not be the soul determinant in any either/or decision about instructional placement, promotion or graduation. Rather, results should be used as indicators of need for early intervention, programmatic changes, or more specific evaluation of learning problems."

The large scale and varied purposive use of standardized tests do not guarantee that these tests are free from deficiencies and apprehensions. Test items lose their credibility due to intensive and repeated use. Their periodical updating on the cost of financial loss, administrative headache and substantial efforts is not easy (Hoffman, 2004). Thorndike (2005) calls these tests expensive as well as time-consuming. Carr and Connie (2004) complaint against their overuse and misuse. Alper, et al. (2001) blame these tests for ignoring diversified characteristics like cultural background, native language and institutional local environment while comparing an examinee's score with his/her age or grade cohort. In the same context, Ormrod (2012) discloses that many people take scores of standardized tests for an indicator of classroom achievement and force teachers to uplift scores on these tests. The teachers feel pressure over such demands covering multi-elaborated while

curriculum in the classroom. Paratore and McCormack (2007) add that teachers have to work in classes and with subject matter while standardized tests are usually long lists of things to be done.

Acknowledging the pros and cons of standardized testing, the researcher decided to work on this direction adopting a more workable approach other than usual analyses. Indeed, traditional item calibration depends upon the ability and numbers of respondents used for the purpose and in this way, results are tentative (Mehrens & Lehmann, 1973). Item difficulty values tend to be positively biased against high ability sample and vice versa. Consequently, item discrimination indices become high for a heterogeneous group but low for homogenous group. This dilemma disturbs the true ability variance. Highlighting more with the same, Stanely and Hopkins (1978) reveal that inequity of norming groups makes the inferences doubted. Hence, achievement tests normed on the same group are credible but Popham (1981) calls such an effort problematic as costs. prompt exhaustion. unmotivated responses and artificiality interpose the task. In the same context, Rasch Measurement Transaction (2006) bears that classical calibration happens imprudent to missing data, adaptive testing, having criterion based pass-fail points and test equating with small groups. Riaz (2008) has discussed various methods of such analysis with all the details and intricacies. It is to clarify that the researcher does not intend to negate traditional methods of test item analysis. Such methods can even out instabilities across test items to yield trusted scores. The need is to meet key criteria pertaining to item development, sampling and sample size, content coverage, validity, textual and linguistic clarity, and test administration etc. to enhance worth of this huge and worldwide enterprise.

In the above scenario, the alternative approach is Rasch calibration. Wright (1967) exposes that Rasch models ascertains object-free instrument calibration and instrument-free object measurement to generalize measurement beyond the specific instrument used. Fluctuation arisen in person abilities due to easy or difficult test is removed from test scores. The same is done to control variation in item difficulties happened due to the responses of a non representative sample. This thing helps to compare objects measured with alike-instruments and to combine or partition instrument to meet new measurement needs, retaining confidence in the procedure. Being a sample-free model, there is no need for big samples and assumes normality of distributions along with its simplicity. Granger (2008) discloses that Rasch calibration estimates a person's probable rating without imputing missing data, test items skipped due to some reason. Further, Rasch modelling accounts for both item-fit and person-fit, while other item response theory (IRT) models address item-fit only. Wisniewski (1992) clears that Rasch model coordinates data to define measurement usefully. It uses the same unit "logit" to quantify item difficulties and person abilities for convenience and meaningful comparisons (Athanasou & Lamprianou, 2009).

Vogel and Engelhard, Jr. (2011) narrate that Rasch model sees the individual learning differences in item-level performance, rather than determining overall group performance as done with statistical procedures like t-test and ANOVA. McArthur (1987) brings forth that Rasch item calibration and goodness-of-fit to the model correspond to classical item analysis procedures but with prominent distinctions. Such distinctions may include the adjustment of difficulty logits of a new test through "linking items" (Popham, 1981), equation of instruments to measure the same trait or performance (Masters & Keeves, 1999), use of Model error estimates for precision of agent and object measurement (Wright, 1978), estimation of probable outcomes of person-item interaction (Salkand, 2007), knowledge of gain scores in the form of meaningful mathematical statements (Wilkerson & Lang, 2007), diagnosis of classroom assessments (Vogel & Engelhard, Jr.,2011), evaluation of item-fit through Rasch measures to avoid computational complexities (Karabatsos, 2000) and also through chi-square fit-statistics to control the suitability of data for the model (Rasch, 1960/1980).

Although, Rasch modelling is quite sensible yet it does not safeguard against all the pitfalls of traditional item analyses. It has its own limiting assumptions. One is unidimensionality. It refers to assess a single ability by a test item that is an ideal case. Its violation happens as the basic cause of misfit of test items (Bond & Fox, 2007). Another assumption is local independence, no impact of response of an item on the response of another item. Baghaie (2008) tells, item-to-item dependence makes parameter estimates biased. Third assumption is the guessing factor, especially working with multiple-choice items (MCIs), where rigorous correct percentage can be achieved just by chance. Being one-parameter logistic model, Rasch model overlooks this reality. Next, Rasch modelling expects from test items to discriminate test takers in a uniform way. This assumption seems hard to defend with finite data when item logistic characteristic-curves deviate from common slopes. Two distinct causes of unequal discriminations are varying reliability of test items and chances of random guessing. Additionally, Rasch modelling also demands some knowledge of and acquaintance with Mathematics in contrast to traditional item analyses.

The limitations of Rasch modelling are irrefutable. No doubt, human and environmental factors disturb test item calibration commonly in all modes used for the purpose. Even then, Rasch model has its own distinctive features. Its capability to work with small samples adequately suits to investigate classroom practices. Elimination of sample effect provides a base for test item selection

varied for purposes. Rasch model makes respondent-to-respondent item-to-item and meaningful the form comparisons in of mathematical statements. For example, an item with a difficulty logit 2.32 is two times harder than an item with difficulty logit 1.16. The same is true of respondents having ability logits 2.32 and 1.16. Further, the use of the common unit "logit" to quantify both difficulty and ability values on the continuum facilitate same latent inter item-respondent comparisons. Hence, an item with a difficulty logit 1.94 is difficult for a respondent with ability logit 1.63 but easy for the one with ability logit 2.19. Additionally, the probability of correct answer on any item by any respondent can be calculated. Similarly, a few items with known difficulty logits can serve as "linking items" to adjust the difficulty of a newly constructed test or equate concurrent tests to appraise same area of performance. Further, Rasch modelling uses the matching of difficulty and ability logits to increase the precision of estimation as an alternative to big samples used to reduce error estimates in traditional calibration methods. The underlying theme of many of such possibilities is the sample-free difficulty measurement and test-free ability measurement. Beyond all these, Rasch model uses simple Mathematics in contrast to other IRT models.

Fit-statistics is an inseparable aspect of Rasch calibration. It sees the extent to which responses meet Rasch assumptions. These assumptions are simple and logical as:

(i) abler persons have greater chances to success on test items than less able persons.

(ii) a person is expected to do better on easy items than hard items.

No doubt, Rasch assumptions provide a justified base for response patterns but practically, these are hard to meet. These assumptions prove quite superficial and demand mechanical responses, putting intervening human and environmental factors aside. Carelessness, fumbling, plodding, response sets, drowsiness, test anxiety, sudden illness, lucky guessing, local climatic conditions and cheating cause bizarre responses. However, Rasch model is a probabilistic model that accounts for inconsistency and awkwardness in measurement, and accepts a plausible magnitude of misfit. It makes room upto  $\pm 0.30$  points fluctuation around the ideal fit of 1.00 due to imperfect human nature in case of MCIs. Items deviating from the prescribed limit 0.70-1.30 may be reconsidered in terms of linguistic, dimensional, discriminative independence, and guessing perspectives to control intervening effects to an affordable extent.

The interpretation of fit-statistics is simple. Values less than 0.70 denote overfit to the model. In simple words, test takers have responded more correctly than the expectations of the model. Overfit is not a very bad thing. Sometimes, inconsistencies of human nature make the responses too good. On the other hand, values more than 1.30 show underfit. It is commonly regarded as misfit. In such a situation, the score on the item is not a valid indicator of its difficulty (Karabatsos, 2000). It reveals that test takers have responded correctly to difficult items but incorrectly to easy ones. This is the real concern and threatens the measurement. One needs worrying about it as extreme misfit cases may mislead the interpretation of the results to reach fake conclusions. Usually, overfit and misfit are reported in terms of outfit and infit mean-squares to denote the nature of randomness along with its magnitude. Outfit tells that mismeasured items were at the extremes of the scale while infit accounts the middle of the scale. Wilkerson and Lang (2007) bring forth an interesting fact that infit is more sensitive to organised mistakes like a string of correct responses or marking the same distracter correct repeatedly. On the other hand, outfit is more distorted by random mistakes like carelessness or test anxiety. In fit perspectives, an unavoidable test item with alarming misfit may be passed through frequent cutting and clipping to make

it "good enough" putting things doing wrong aside. Kreiner (2012) tells that George Rasch was very cautious about such a bad happening and asked to manage a remedy for it. However, a few content-wise representative items may also be included in the test without considering fit dilemma. Experts have no objection to such rare items in the testing scheme. It is noteworthy that the fit criteria are same for items and respondents but from a substantive view, items and respondents differ in this regard. Items are supposed to do betterly than respondents. Hence, fit rules are applied more strictly to items than to respondents. If a few respondents do not interact items accordingly, no need worrying. The same is not true with test items. Karabatsos (2000) suggests a simple way of fit-analysis through Rasch measures. This method characterises response-fit on the same scale as used for difficulty and ability logits. It uses the unit "logit noise" to describe misfit. This method considers absolute values of unwanted responses to quantify misfit. The average misfit logit can be calculated for each person and item, a group of persons or a set of items and even for the entire sample or the instrument.

For the current study, MCIs were used. No doubt, much admissible criticism on this format is evident in literature after the emergence of new trends in assessment. Even then, one has to rely on MCIs for a number of reasons. A talented test item writer can use this format to appraise comprehension, interpretation, application, analysis or synthesis to arrive at the keyed answer. These tests also seem beneficial to make judgments, inferences and generalizations. Popham (1981) says that MCIs adequately cover advanced intellectual skills and significant attitudinal dispositions besides testing factual knowledge. Athanasou and Lamprianou (2009) reveal that MCI format is used for large scale examinations such as the Higher School Certificate and the Scholastic Aptitude Test, the mathematics and science competitions in primary and high schools developed by the University of New South Wales, the Basic

Skills Tests used in many states, educational tests distributed by the Australian Council for Educational Research and Overseas Skills Recognition Tests due to their versatile coverage of subject areas and adoptability. Their utility is at the same level in many countries other than Australia. Especially, countries like UK and USA make copious use of this format. In Pakistan, Federal and Provincial Public Service Commissions and other recruitment agencies use this test format for appraisal. Entry tests in medical and engineering are solely MCI based. The National Testing Service (NTS) adopts this format all alone to judge candidacy for admission to universities of Pakistan. These tests also make a major part of assessment at school and college level.

The above stated virtues and functions of MCI format are not a guarantee for their perfection as assessment tool. These are equally criticised for a number of reasons. MCIs seem helpless to bring fluency and spontaneity in written expression and retard originality. This format does not provide opportunity to synthesize thoughts and write out creative solutions. People blame MCIs for common mistakes in spellings and sentence structures by students. Simply, recognition of a response is inferior to its construction. Problems like ambiguity in stems, complicated syntax, heterogeneous options, unintentional clues for correct answer, grammatical inconsistencies between stem and options, unfair positioning of correct answer, guessing and other concerns are common with MCIs especially when the test developer is inexperienced or non-professional. Hence, a judicious and blended use of MCIs can sufficiently justify their presence in modern assessment. Woolfolk(2011) asks for the use of constructed-response format, authentic assessment and portfolio assessment to deal with testing problems common with MCIs. In the same way, Kaplan and Saccuzzo (2007) prefer written tests and work samples to usual MCIs to assess students. However, problem of comparison of performance with these methods awaits future work. Haladyna and

Rodriguez (2013) introduce some modified and improved versions of MCIs like multiple-mark item format, "uncued" multiple-choice, ordered multiplechoice and two-tiered diagnostic multiple-choice to compensate existing flaws of this format.

The researcher opted MCIs for the study due to the reasons like Pakistani students' acquaintance with this format, their focal weightage in assessment and after all, data need to be based on dichotomous choice for test item calibration though Simple Rasch Model. However, their exists Rasch's Partial Credit Model to accommodate partial award as in the case of short answer questions but all this is beyond the scope of current study. The researcher wants to share another compulsion with the readers that the test items of this study mostly cover recall of factual knowledge and comprehension. These overlook advanced cognitive levels. All this reflects local standards and the researcher is helpless in this regards. Question papers of boards of intermediate and secondary education (BISEs) seem to revolve around initial two phases of cognitive domain although; much has been written in the syllabus about objectives of teaching Mathematics at secondary level.

## Method

The study was carried out as below.

## **Participants**

The total 59,168 students of grade 9 admitted to 718 secondary schools for session 2010-11 in the jurisdiction of Bahawalpur division comprised the population of the study. Further, 405 boys and 313 girls secondary schools were classified locality-wise and ownership-wise to ensure the presence of categorical subgroups of interest in the sample. The cluster sampling was used to select randomly 3 schools from each category. All the students of grade 9 of these schools were included in the sample. The categorical breakdown of the sample is given in the table.

Sex	-wise	Resi	dence-wise	Institution-wise		
Repres	entation	Repi	resentation	Representation		
Males	Females	Rurals Urbans		Public Schools	Private Schools	
333	309	326         316           (50.78%)         (49.22%)		345	296	
(51.87%)	(48.13%)			(53.74%)	(46.26%)	
Total 642		Т	otal 642	Total 642		

**Breakdown of Sample** 

### Instrument

Two test forms entitled test form A and test form B equivalent in content and format, each bearing 45 items were developed from the prescribed textbook of Mathematics for grade 9, retaining a recommended proportion among various content areas and focusing instructional objectives. During pilot testing, two items from each test form were debarred owing to textual and structural complexities. To compensate curtailing, some items were replaced to retain equivalence in test forms. Some other items were rephrased for clarity. The final draft of each test form comprised 43 items (see appendix). In this way, the intended test to be semi-standardized consisted of 86 items. The test was administered in the form of two separately-timed tests. The reliability coefficient remained 0.92 when calculated with alternate-forms method

## Procedure

The heads of the sampled schools were consulted in advance to seek permission for the administration of componential tests with the assistance of concerned class teachers. Each time and elsewhere, the sub tests were got attempted in 3<sup>rd</sup> and 4<sup>th</sup> periods in two consecutive days. The time allowed was one hour and twenty minutes, out of which 20 minutes were specified for instruction and seating arrangements, while 1 hour was for filling the profile page and attempting the test. Efforts

Teachers of Mathematics and researchers were consulted to incorporate their judgment for the validation of the tool. Certain arguments, claims and previous evidences were considered in this regard in a panel discussion. Replacing and rephrasing of items in the initial draft of the tool as suggested by experts were made to cover defined objectives for the teaching of Mathematics and ensure due weightage among various componential content areas of the prescribed text as Sets 18 items, Systems of Real Numbers 8 items, Logarithms 12 items, Algebraic Expressions and Factorization 18 items, Matrices and Determinants 12 items, and Geometry 18 items. Each test item of the tool with its difficulty logit (di), Model error estimate (SE), average logit noise (|K|), infit mean-square (vi) and outfit mean-square (ui) respectively is as under.

were made to keep instructions uniform and orderly. A usual but fair and conducive environment was ensured to secure the external validity of the test.

#### **Data Analysis and Results**

In this study, difficulty logits for 86 items ranged from-1.40 to 0.89 with a mean 0.003 as shown in table 1. Item 68 with difficulty logit-1.40 happened the easiest. It was respondent correctly by 529 students out of 642. The item 16 was the hardest with a logit value 0.89 which was marked correct only by 246 respondents out of 642. Difficulty logits for the remaining 84 items were found between these two extremes. It was concluded that all items were moderate in terms of difficulty. No need was felt to exclude any item from the test owing to its difficulty or ease.

In contrast to item difficulties, ability logits were more dispersed on both sides with values-1.67 to 4.83 around a mean 0.06. The minimum earned score by any respondent was 15 and the maximum was 85 out of total score 86. The distribution of respondents' earned scores on the test was almost normally distributed. It means that the sample was adequately representative.

Model error estimates determine the accuracy of measurement. Their low values indicate exactness and accuracy in item difficulty and respondents' ability measurement. In current study, error estimates for items remained low and consistent with the mean value 0.09. It means that difficulty logits were measured more precisely and sharply. Two things cause this precision. One was the appraisal of each test item by 642 respondents with varied range of abilities due to a big and representative sample. The other was the matching between difficulty and ability logits due to moderate nature of many test items.

The average Model error estimates remained high i.e., 0.29 for ability measures. It means that respondents' abilities were not measured as precisely as item difficulties. The cause behind the phenomenon was that each respondent encountered only 86 items, with relatively condensed range of difficulties.

The "logit noise" is a simple measure of fit-statistics introduced by Karabatsos (2000). It uses the same scale to calculate misfit as used for difficulty and ability logits. A positive value denotes that response is above the expectation of the model and vice versa. A zero reflects fit of response with the model.

The average logit noise remained 0.18 for items and 0.27 for respondents in current study as shown in table 1. It means that there was less distortion in difficulty measurement than ability measurement. Hence, difficulty logits were measured more precisely. It is notable that all the determinants of Model error estimates for difficulties and abilities work alike in the quantification of logit noise values due to the use of same scale "logits" and the same latent continuum for these. That is why; this method is termed as "Fit-analysis through Rasch Measures." This method has its own limitations. It is hard to say that the average logit value 0.18 for items and 0.27 for respondents are good or bad indicators of the fit phenomena because no unanimous norms are available to decide about acceptable distortion. Further, overfits (positive values) or underfits (negative values) are jumbled due to taking both values positive in the calculation of average "absolute value" of logit noise (misfit). The average logit noise can be calculated for an item or the entire test or in turn, for a respondent or the entire sample. Owing to mentioned deficiencies, chi-square fit-statistics is much better than this method.

Table 1

Summary Statistics for Test Items and Respondents

Sr No	Statistics Aspects	No. of Cases	Ra	Range of Values			
SI. INO.	Statistics Aspects	No. of Cases	Minimum	Maximum	Mean Value		
1 Difficu	lty Logits (di) Items	86	-1.40	0.89	0.003		
2 Ability	v Logits (bv) Persons	642	-1.67	4.83	0.06		
3 Model	Error Estimates (SE) Items	86	0.09	0.12	0.09		
4 Model	Error Estimates (SE) Persons	642	0.23	1.05	0.29		
5 Averag	ge Logit Noise  K  Items	86	0.06	0.37	0.18		
6 Averag	ge Logit Noise  K  Persons	642	0.23	0.34	0.27		
7 Infit M	lean Square (vi) Items	86	0.74	1.12	0.93		
8 Outfit	Mean Square (ui) Items	86	0.56	1.22	0.92		
9 Infit M	lean Square (vi) Persons	642	0.93	1.19	1.05		
10 Outfit	Mean Square (ui) Persons	642	0.49	1.62	1.08		

In chi-square fit-statistics method, the ideal index for average infit and outfit mean-squares is 1.00. However, fluctuation with a range 0.70-1.30 is admissible around the ideal fit 1.00. The criteria are same both for item and person measures. The "infit" accounts for the middle of the difficulty or ability scale while "outfit" covers the extremes of these scales.

In this study, the average infit mean-square was 0.93 for items as shown in table 1. The individual test item measures also remained within the prescribed limit 0.70-1.30. Hence, no overfit or misfit was observed. Simply, middle of the difficulty scale met Rasch assumptions adequately.

The case was not the same with outfit mean-squares. The average outfit mean square remained 0.92. Three individual test items 2,11 and 60 violate the Rasch assumptions with an average

value of 0.60 falling below 0.70, the lower admissible limit as shown in table 2. It indicated an overfit i.e., the responses on these items were more correct than the usual at the extremes of the test scale.

Again in current study, the average infit mean-square was 1.05 for respondents. The individual respondents' ability measures also remained within the prescribed limit 0.70-1.30. Hence, no overfit or misfit was observed. Simply, middle of the ability scale met Rasch assumptions sufficiently.

The case was not same with outfit mean-squares. The average outfit mean-square remained 1.08. Individually, 8 males and 18 females with average overfit 0.65 and 0.62 respectively falling below 0.70, the admissible lower limit as shown in table 2. It means, responses of these respondents were more correct than the usual at the extreme of the scale. Similarly, 8 males and 11 females with average misfit 1.52 and 1.48 respectively moved up 1.30, the upper admissible limit. It means that these respondents answered correctly on difficult items and incorrectly on easy items at the extremes of test scale. It is noteworthy that 4 females commonly shared overfit and misfit but no male behaved in this manner. Individually, there were more overfit cases with females and misfit (underfit) cases with males The Rasch analysis with model-fit statistics revealed an overall good correspondence between the observed and expected scores based on the model. The number of mismeasured cases and their intensity was tolerable to declare fit-statistics good to include all the items in the intended semi-standardized test.

## Table 2

						Individual	
Cases		Mean Values					
Aspect	Mean Squ	ares Tot	tal Cases				
						Overfit	
	Mi	sfit	Overfit	Mis	sfit		
Itoms	Infit	-	-	-	-	-	
nems	Outfit	3	3	-	0.60	-	
					0.00		
	Infit	-	-	-	-	-	
Dersons	Outfit	*16	6	10	0.65	1.52	
T CISONS	Infit	-	-	-	-	-	
	Outfit	**23	18	9	0.62	1.48	

Mismeasured Cases Beyond Plausible Limit 0.70-1.30

\*Males \*\*Females

The graphical illustration of selected item and person measures and their fit with the model is shown in figure to represent fit phenomenon more clearly.



## **Discussion and Recommendations**

The researcher describes a quantitative approach through the lens of Rasch measurement theory for developing a semi-standardized test in Mathematics at grade 9. A secondary purpose was categorical comparison of respondents' performance on the test. No significant difference of performance was seen gender-wise and locality-wise. In this way, the substantive results of the study remained consistent with Ahmad (2004) and Ali (2008) who concluded the very same from their studies. The "norms" of a standardized test give meanings to raw scores of subsequent test takers. For this, respondents' raw scores on the test were transformed to percentile ranks and T scores in a table annexed with the study. The mean, standard deviation and standard error of measurement of this semi-standardized test remained 48.92, 14.63 and 0.577 respectively to serve as standard frame of reference for comparison.

The underlying themes of Rasch modelling reveal that it is considerable to address

measurement problems faced in Pakistan. At present, boards of intermediate and secondary education (BISEs) do not equalize difficulty level of their question papers in different subjects for intermediate examinations. secondary and Consequently, the question paper in a particular subject in one board may be comparatively difficult, moderately difficult in other board and considerably easy in another board. The same is true of question papers of group I and group II in the same board in subjects with heavy enrolment. Owing to this, examinees suffer in two ways. Overall positions in each board are severely disturbed. For example, overall first position in the board of a pre-engineering student may be scored due to easiness of question paper in Mathematics than that of Biology of a medical student. The phenomenon is more confused for position holders of general science and humanities where elective subjects are so many. Luck seems to favour particular students due to appearing in particular subjects. Secondly, merit for medical, engineering other competitive programs and becomes

ridiculous when candidates passing from different boards are rated at the face value of their marks overlooking the passing board and difficulty of its question papers. Owing to this, many admission seekers are deprived of their right while others are favoured unintentionally. All this demands that there is need to recognise merit prior to ascertain it.

The situation is not different with recruitment through Punjab Public Service Commission. Numerous candidates apply for some posts. Constraints of infrastructure compel administration to conduct the written test in two groups for the same post. Each group is assessed with a different paper on MCI format. No conscious effort is made to adjust the difficulty level of intended tests. Luck or hardluck is greatly tied to appear in a particular group. The affected candidates feel this controversy and even unfold it informally but all in vain. Further, in Pakistan, the National Testing Service (NTS) has been assigned the responsibility to conduct eligibility test for admission to M.A., M.Sc., M.Phil. and Ph.D. programs of various universities. The test results remain valid for two years. The NTS arranges four test yearly for the same purpose. In this way eight tests are administered in two years. Undoubtedly, these tests vary in difficulty. The claim is evident from varied corresponding percentile ranks of same raw scores on different tests, assuming the normality of distributions due to heavy participation each time. The universities use raw scores of these tests to determine merit for admission. Once again, fortune or misfortune works. Candidates passing from separately-timed tests may reap the benefit of appearing in a particular test and vice versa.

The above scenario reflects need for fair and meritorious attempts. First of all, acknowledgement of the bitterness of current practices is necessary to move ahead. Next steps might be assigning the critical task of item writing to professionals instead of laymen, reconsideration of previous test items to note flaws to have insight in subsequent item writing, panel discussion on written items to refine and enrich these, proportional representation of cognitive levels in item writing, preparation of quality booklets bearing principles and practical examples of test item writing in various subjects, holding workshops for the orientation of the test developers, preparation of item banks for different purposes and thoughtful allocation of final items in terms of difficulty and cognitive level to all intended question papers. Besides these technical tips, human factors like sincerity, enthusiasm, devotion and sense of responsibility can add to this art. All the above points are worthy of consideration but there lies one thing more. This is the empirical touch to test items. It will reveal the appropriateness of our claims and logic for the task of judgmental calibration. The researcher means test equating with the help of "linking items" and fit trial of constituting items.

Masters and Keeves (1999) state three methods of test equating. One is "common item threshold difference method." In this method, two sets of items with purposely chosen sufficient common items are separately calibrated. The average threshold difference between common items of two tests is calculated as the equating constant between the two scales. The common items between two sets also assist to know error estimates with mean difference. The second method is called "anchor item equating method." In this method, one set of items is calibrated to form a scale. The threshold values of particular items in the calibrated set which are same to other set are used in the calibration of the second set of items. Another, computer-based readily applied method is "concurrent equating method." It combines many data sets with common items in a way that a single calibration of all data sets is undertaken. Owing to the large storage capacity of computer, intended data sets are brought to a common scale. The procedure ascertains more consistent and stronger

measures of items in various equated data sets (Baker & Al-Karni, 1993).

In Pakistan, each BISE may use one of the first two methods to equate its question papers of first group and second group in the subjects with heavy enrolment at secondary and intermediate level. The Punjab Public Service Commission may repeat the same practice for its tests in two groups against the same post. The third method is feasible for inter-boards equivalence of question papers in the same subject where at least ten papers are needed. It is equally beneficial for NTS to equal all the intended tests for candidates seeking admission to the same program. The researcher suggests four precautions to work with these methods. These are scatter positioning of "linking items" in the test, coverage of same cognitive levels in common items as in the total test, forestalling the occurrence of probable breaches of test security for two separately-timed tests and restrain students from cheating, a bitter experience in Pakistan.

The major recommendations emerged from the study were its replication, its conduction at other times during the session to examine possible differences, province wide and countrywide expansion of the study to set respective level norms, diagnosing the causes of overfit or underfit for items and persons, suggesting BISEs to use gratis data for research and calibration, and using advanced software for intensive, excessive and speedy data analysis. Prospective researchers are suggested to expand the use of Rasch measurement theory to calibrate test items in different disciplines for reliable results. Further, the need is to make Rasch's Partial Credit Model customary to calibrate short answer questions.

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# Appendix

						]	<b>Fest</b> F	orm A	<u>L</u>		
Su	ıbject: Ma	athematic	es Lev	el: Grad	le 9	No.			of	Items:	43
Ti	me:1 Hou	ır									
No	ote: Each	item bea	ars for	ur respo	onses a	,b,c and	d. En	circle	the correct	one. Please note that omi	tted items
wi	ill be trea	ted as inc	correct	t.							
Se	<u>ets</u>	d	i,	SE,	/K/,	vi,	и	i			
1	The not	ation of t	he set	of integ	ers is			(0.38	8, 0.09, 0.18	, 0.86, 0.78)	
	(a)	E	(b)	Р	(c)	0	(d)	Ζ			
2	Null set	is denote	ed with	h the sy	mbol			(-0.4	8, 0.10, 0.09	9, 0.74, 0.59)	
	(a)	{ }		(b) {	0 }	(c) 0		(d)	{ }		
3	If the nu	umber of	eleme	nts in a	set A a	tre "n" th	nen the	e numl	per of eleme	ents in P (A) will be	
	(0.2	24, 0.09,	0.15, (	0.85, 0.8	86)						
	(a)	$n^2$	(b)	2 <sup>n</sup>	(c)	$2^{2n}$	(d)	2n			
4	Tabular	form of t	he set	${x/x N}$	^10 <x< td=""><td>&lt;15} is</td><td></td><td></td><td>(0.10, 0</td><td>0.09, 0.22, 0.77, 0.82)</td><td></td></x<>	<15} is			(0.10, 0	0.09, 0.22, 0.77, 0.82)	
	(a)	{10,11	,12,13	,14}		(b)	) {11	,12,13	3,14,15}		
	(c)	{11,12	,13,14	.}		(d)	) {1(	),11,12	2,13,14,15}		
5	A set ca	n be pres	ented	through	l			(-0.8	7, 0.10, 0.00	5, 1.02, 1.07)	
	(a)	Descrij	ptive I	Method		(b)	) Tał	oular H	Form		
	(c)	Set Bu	ilder N	Notation	l	(d)	) All	the al	bove metho	ds	
6	The	e point (2	,4) lie	s in the		quadr	ant of	Cartes	sian co-ordi	nate system.	
	(-0.	.14, 0.09,	0.15,	0.92, 0.	.85)						
	(a)	$1^{st}$	(b)	$2^{nd}$	(c)	3 <sup>rd</sup>	(d)	$4^{\text{th}}$			
7	The	e number	of ele	ements i	n the d	lomain s	et of th	ne bin	ary relations	$\{(1,1),(2,-1),(2,-3)\}$ is	
(0.	.47, 0.09,	0.20,0.7	6, 1.10	0)							
	(a)	2	(b)	3	(c)	4	(d)	6			
8	The syn	nbol '^' st	ands f	or				(-0.5	2, 0.10, 0.2	1, 0.82, 0.94)	
	(a)	or	(b)	and	(c)	less the	an (d)	grea	ter than		
9	Sets are	usually c	lenote	ed by		_ alphab	ets.		(-1.05,	0.11, 0.16, 0.93, 0.87)	
	(a)	Latin	(b)	Greek	(c)	Englis	h (d)	Aral	bic		
						<u>System</u>	s of R	eal Ni	umbers		
10	) Multipli	icative in	verse	of $\frac{1}{a}$ is				(0.29	9, 0.09, 0.13	, 0.90, 0.92)	
	(a)	$\frac{1}{a}$	(b)	$\frac{1}{a}$	(c)	-a	(d)	a			
11	$(4)^3$ ec	quals to					(-0.	08, 0.0	09, 0.28, 0.7	79, 0.64)	
	(a)	$4^{5}$	(b)	$4^{6}$	(c)	4 <sup>8</sup>	(d)	49			
12	$X^{2/3}$ can	be writte	en as				(0.3	8, 0.0	9, 0.19, 0.82	7, 0.80)	
	(a)	$\sqrt{X^{3/2}}$	(b)	$2\sqrt{X^3}$	(c)	$3\sqrt{X^2}$	(d)	$(\sqrt{X})$	$\frac{3}{2}$		
13	8 If x=4-	$\overline{15}$ , then	the v	alue of	$\frac{1}{x}$ is				(0.60, 0.09,	0.20, 0.86, 1.12)	

(a)  $1/4-\sqrt{15}$  (b)  $1/4+\sqrt{15}$  (c)  $-4+\sqrt{15}$  (d)  $4+\sqrt{15}$ **Logarithms** 14 Antilogarithm table was developed by (-0.64, 0.10, 0.12, 0.95, 0.84) (b) Henry Briggs (c) John Napier (d) Al Khwarizmi (a) Jobst Burgi 15 The common form of  $8.24 \times 10^{-4}$  is (0.70, 0.09, 0.17, 1.06, 1.06)(a) 82400 (b) 0.000824 (c) 0.0000824 (d) 8240000 16 The logarithmic form of  $4^3$ =64 is (0.89, 0.09, 0.37, 1.12, 1.15) (a)  $\log_3 64=4$  (b)  $\log_3 4=64$ (c)  $\log_4 64=3$ (d)  $\log_4 3=64$ (-0.33, 0.09, 0.16, 0.94, 1.13) 17 The characteristics of log 325 is (a) 0 (b) 1 (c) 2 (d) 3 18 The single logarithmic form of log 5+log6-log2 is (0.47, 0.09, 0.16, 0.92, 0.92)(a)  $\log \frac{5\pi 6}{2}$  (b)  $\log 5+6-2$  (c)  $\log 15$  (d)  $\log 9$ 19 If  $\log x=3.6862$ , then its antilog 4855 with a right decimal point is (0.56, 0.09, 0.24, 0.79, 0.81) (b) 485.5 (c) 0.004855 (d) 0.0004855 (a) 48.55 **Algebraic Expressions and Factorization** 20 The co-efficient in  $3x^2$  is (0.24, 0.09, 0.15, 0.96, 0.90) (a) 3 (b) x (c) 3x (d) 2 21 \_\_\_\_\_ is a polynomial expression. (0.60, 0.09, 0.22, 1.09, 1.06)  $\frac{-2}{x^3}$  (b)  $6x^{-2}$  (c)  $-x^3$  (d)  $3x^{-2}+\frac{1}{x^3}+2$ (a) 22 Polynomial expressions can be classified with respect to (0.01, 0.09, 0.20, 1.06, 0.87) (a) terms (b) variables (c) degrees (d) all the three 23 In algebraic expressions, ascending/descending order is made with respect to (0.33, 0.09, 0.26, 0.99, 1.08) (a) co-efficients (b) exponents (c) variables(d) something else 24  $(a+b)^2+(a-b)^2$  equals to (0.24, 0.09, 0.17, 0.90, 0.88) (b)  $2(a^2 + b^2)$  (c)  $a^4 - b^4$  (d)  $(a+b)^4$ (a) 4ab 25 In  $x^2$ -5 x +6/ x -2, x -2 is (0.01, 0.09, 0.17, 0.94, 0.82)(a) dividend(b) divisor (c) quotient (d) remainder 26 If x + y=2 and xy = 3, find the value of  $x^2 + y^2$ (0.70, 0.09, 0.28, 1.08, 0.90)(d) -2 (a) 10 (b) 6 (c) 2 27 Factorize  $x^{2}+x-6$ (0.24, 0.09, 0.14, 0.98, 0.93) (a) (x+3)(x+3)(b)(x-3)(x-3)(c)(x+3)(x-3) (d) (x-2)(x+3)28 Methods to find Least Common Multiple (LCM) are (-1.40, 0.12, 0.14, 0.89, 0.96)(c) 4 (d) 5 (a) 2 (b) 3 Matrices and Determinants 29 Matrices were introduced by Mathematician (-0.08, 0.09, 0.09, 0.95, 0.79) (a) Arther Kelley (b) Cramer (c) Demorgan (d) Euclid 30 A matrix with different number of rows and columns is called a \_\_\_\_\_ matrix (-0.74, 0.10, 0.18, 1.01, 1.08) (b) column (c) rectangular (d) square (a) row

31	$\begin{array}{cccccccccccccccccccccccccccccccccccc$	
	(a) Zero (b) diagonal(c) scalar (d) identity	
32	is possible on $[xy]^{-1}$ (0.38, 0.09, 0.20, 0.79, 0.78)	
52	2 (0.56, 0.09, 0.20, 0.79, 0.76)	
	(a) Addition(b) Subtraction (c) Multiplication (d) Division	
33	If $A = \begin{pmatrix} 6 & 4 \\ 1 & 3 \end{pmatrix}$ and $B = \begin{pmatrix} 6 & 1 \\ 4 & 3 \end{pmatrix}$ then B is a/anmatrix of A. (0.65, 0.09, 0.13, 0.95, 0.86)	)
	(a) ad joint (b) transpose (c) singular (d) inverse	
34	The matrix resulted from the multiplication of $\begin{bmatrix} 3 & 1 & 1 & 2 \\ 1 & 1 & 1 & 1 \end{bmatrix}$ is (0.19, 0.09, 0.28, 1.10, 0.93)	
	(a) $\begin{array}{cccccccccccccccccccccccccccccccccccc$	
	Geometry	
35	The word "Geometry" has been derived from       (-0.70, 0.10, 0.15, 0.92, 1.15)	
	(a) Latin (b) Greek (c) Hebrew (d) Arabic	
36	Literally, the word "Geometry" means the measurement of (-0.70, 0.10, 0.10, 0.95, 0.86)	
	(a) earth (b) rocks (c) ocean (d) air	
37	A fundamental agreement related to almost all the branches of Mathematics is called a/an	
(0.5	56, 0.09, 0.22, 1.05, 1.08)	
•	(a) axiom (b) postulate (c) corollary(d) rider	
38	The second part (to be proved) of the statement of a geometrical theorem starts with the word	
	(0.74, 0.09, 0.36, 1.07, 0.88)	
20	(a) if (b) but (c) and (d) then $(0.50, 0.10, 0.15, 0.00, 0.02)$	
39	The total of supplementary angles is $(-0.58, 0.10, 0.15, 0.90, 0.82)$	
10	(a) $90^{\circ}$ (b) $120^{\circ}$ (c) $180^{\circ}$ (d) $360^{\circ}$	
40	Out of six basic elements of a triangle, at leastare necessary to be known to construct	а
	triangle. $(0.55, 0.00, 0.00, 0.07, 0.77)$	
	(0.50, 0.09, 0.08, 0.97, 0.77)	
41	(a) $2$ (b) $3$ (c) $4$ (d) $5$	
41	In a right angled triangle, the opposite side of $50^{\circ}$ , is the of its hypotenuse in measure.	
(a)	(0.19, 0.09, 0.23, 0.78, 0.92)	
(a)	The sum of the measures of the three angles of a triangle is $(-0.91, 0.10, 0.17, 0.86, 0.87)$	
72	(a) $120^{\circ}$ (b) $135^{\circ}$ (c) $180^{\circ}$ (d) $225^{\circ}$	
43	In $a/an$ triangle all the three right bisectors of sides are concurrent at the midpoint of the	าค
15	hypotenuse	10
	(0.65, 0.09, 0.33, 1.05, 0.93)	
(a)	acute angled (b) right angled (c) obtuse angled (d) equilateral	
. /		
	<u>Test Form B</u>	

Subject: Mathematics Level: Grade 9No.ofItems:43

## Time:1 Hour

Note: Each item bears four responses a,b,c and d. Encircle the correct one. Please note that omitted items will be treated as incorrect.

Sets {0,1,2,3,.....} is a set of.....numbers. 1/44(0.10, 0.09, 0.13, 0.98, 1.06)/(a) whole (b) natural (c) prime (d) real 2/45(-0.33, 0.09, 0.11, 0.98, 0.85) Symbol stands for (a) proper set. (b) equal set. (c) sub set. (d) universal set. 3/46 If the number of elements in a set A is 'm' and in a set B is 'n', then the number of elements in AxB will be (0.15, 0.09, 0.18, 0.96, 0.85) (b)  $2^{m-n}$ (c)  $2^{m \times n}$ (d)  $2^{m \div n}$ (a)  $2^{m+n}$ 4/47 The set building notation of the set  $\{2,3,5,7,11\}$  is (-0.03, 0.09, 0.28, 0.86, 1.12) (b)  $\{x/x P 2 x 7\}$  (c)  $\{x/x P 3 x 11\}$  (d)  $\{x/x P 2 x 11\}$ (a)  $\{x/x P 3 x 7\}$ (-1.05, 0.11, 0.11, 0.93, 0.75) 5/48 The methods to present a set are (a) 2 (b) 3 (d) 5 (c) 4 6/49 Point \_\_\_\_\_\_ lies in the third quadrant of Cartesian co-ordinate system. (0.10, 0.09, 0.18, 0.95, 0.84) (a) (2,3) (b) (-2,3) (c) (2,-3) (d) (-2,-3)7/50 The number of elements in the range set of the binary relations  $\{(1,-1), (2,1$ (2,-3)} is (0.24, 0.09, 0.20, 1.03, 0.88) (a) 2 (b) 3 (c) 4 (d) 6 8/51 The symbol 'U' denotes \_\_\_\_\_ \_\_\_\_ set. (-1.12, 0.11, 0.12, 0.92, 0.76) (a) union (b) intersection (c) subset (d) universal set (-0.23, 0.09, 0.14, 0.95, 0.86) 9/52 The right presentation of a set is (a)  $A = \{a,b,c,d\}$  (b)  $a = \{a,b,c,d\}$  (c)  $A = \{A,B,C,D\}$ (d)  $a=\{A,B,C,D\}$ Systems of Real Number 10/53 The set has no closure property with respect to addition and multiplication (-0.14, 0.09, 0.09, 0.84, 0.99) (a)  $\{0\}$ (b)  $\{0,1\}$ (c)  $\{1,0\}$ (d)  $\{0,-1\}$ 11/54  $3x10^{\circ}$  is equal to (0.19, 0.09, 0.15, 0.89, 0.91)(a) 0 (d) 300 (b) 3 (c) 30 12/55  $\sqrt{x} \times \sqrt{x}$  equal to (0.33, 0.09, 0.18, 1.00, 1.00) (a)  $x^2$ (b) 2x (c) x (d)  $\sqrt{2x}$ 13/56 If  $x = 2 + \sqrt{3}$ , then the value of  $x + \frac{1}{x}$  is (0.47, 0.09, 0.14, 0.77, 0.82)(a)  $2 - \sqrt{3}$ (c)  $2\sqrt{3}$ (b) 4 (d) 2x **Logarithms** 14/57 The base of common logarithm is (-0.70, 0.10, 0.15, 0.86, 0.78) (a) e (b) 2 (c) 5 (d) 10 15/58 The scientific notation of 99.99 is (0.24, 0.09, 0.24, 0.94, 0.90)(b)  $9.999 \times 10^{-1}$ (c)  $9.999x10^3$  (d)  $9.999x10^{-3}$ (a) 9.999x10

16/59	The value of x for $\log_2 x=5$ is (0.51, 0.09, 0.21, 0.99, 1.06)
	(a) 7 (b) 10 (c) 25 (d) 32
17/60	The logarithmic statement with characteristics -2 is (-0.14, 0.09, 0.09, 0.83, 0.56)
	(a) log25 (b) log 0.07835 (c) log 400.3 (d) 0.00329
18/61	If $\log 3 = 0.4771$ and $\log 5 = 0.6990$ , then the value of $\log 5\sqrt{3}$ is (0.65, 0.09, 0.28, 1.04, 0.98)
	(a) 0.5881 (b) 1.1595 (c) 0.4604 (d) 0.1110
19/62	The statement with the greatest antilog value is $(0.47, 0.09, 0.14, 0.86, 0.96)$
	(a) log 3.2201 (b) log 5.4654 (c) log 0.9837 (d) log 0.09999
	Algebraic Expressions and Factorization
20/63	Polynomial expression $3x^2+4y+6z+1$ bears variables. (-0.38, 0.09, 0.32, 0.79, 1.15)
	(a) 2 (b) 3 (c) 4 (d) 7
21/64	$5x + \frac{3}{2}$ is polynomial with co-efficient as numbers $(0.42, 0.09, 0.23, 0.92, 0.93)$
21/01	$\frac{1}{4}$ is polynomial with co efficient as <u></u> numbers. (0.12, 0.05, 0.25, 0.52)
	(a) natural (b) rational (c) irrational (d) integral
22/65	The degree of the polynomial expression $5x^3yz^5$ is (0.01, 0.09, 0.18, 0.91, 0.86)
	(a) 5 (b) 8 (c) 9 (d) 10
23/66	The ascending order of algebraic expression $y^4$ -6-4y+ $\frac{12}{32}$ + $\frac{9}{34}$ + $\frac{4}{32}$ is (0.29, 0.09, 0.27, 1.07, 1.17)
	33 34 32
	(a) $\frac{9}{14} + \frac{12}{12} + \frac{4}{12} - 6 - 4y + y^4$ (b) $y^4 - 4y - 6 + \frac{4}{12} + \frac{12}{12} + \frac{9}{14}$
	yr y5 y2 y2 y3 yr
	(c) $-6-4y+y^4+\frac{4}{y^2}+\frac{12}{y^3}+\frac{9}{y^4}$ (d) $\frac{4}{x^2}+\frac{12}{y^3}+\frac{9}{y^4}+y^4-4y-6$
21/67	$(a+b)(a^2 - ab+b^2)$ equals to $(0.10, 0.00, 0.21, 0.07, 0.82)$
24/07	(a+b)(a - ab+b) equals to $(0.19, 0.09, 0.21, 0.97, 0.82)$
25/68	(a) $(a+b)$ (b) $(a-b)$ (c) $a = b$ (d) $a + b$ Solve $3x^2y^2 + 3xy^3$ (1.40, 0.12, 0.13, 1.00, 0.05)
23/08	Solve $3x y = 5xy$ (-1.40, 0.12, 0.13, 1.00, 0.55) (a) $x/y$ (b) $y/y$ (c) $x^3y^5$ (d) $1/x^3y^5$
26/69	Find the value of 'n' to make $x^2 \pm x_{-}n$ a complete square (0.51, 0.09, 0.26, 0.84, 0.87)
20/07	(0.57, 0.07, 0.20, 0.07) (a) $1/4$ (b) $1/8$ (c) $-1/4$ (d) $-1/8$
27/70	Factorize $x^3+125$ (0.10, 0.09, 0.13, 0.95, 0.85)
(a) (x-	( $x^{2}-5x+25$ ) (b) ( $x^{2}+5x+25$ )
(c) (x-	(4) (x <sup>2</sup> +5x+25) (d) (x <sup>2</sup> -5x-25)
28/71	If AxB=LxH, then A will be equal to $(-0.98, 0.09, 0.15, 1.02, 0.94)$
	(a) $B/LxH$ (b) $LxH/B$ (c) $BxH/L$ (d) $BxL/H$
	Matrices and Determinants
29/72	Matrices are typically used in (-0.14, 0.09, 0.07, 0.99, 1.02)
	(a) Engineering and Medical. (b) Mathematics and Physics.
	(c) Economics and Biology. (d) Statistics and Chemistry.
30/73	Two matrices are said to be equal if there are equal. (0.10, 0.09, 0.18, 0.95, 1.21)
	(a) rows (b) columns (c) corresponding elements (d) all the three
21/74	2 0 = 0
51/74	$0  2  18 \text{ a} \qquad \qquad 113 \text{ a} \qquad (-0.42, 0.09, 0.10, 0.89, 0.87)$
	(a) scalar (b) diagonal(c) zero (d) identical

32/75	Which operations are possible on two matrices of the same order? (0.29, 0.09, 0.36, 1.09, 0.94)							
	(a) Addition and subtraction (b) Addition and multiplication							
	(c) Subtraction and multiplication (d) All mentioned operations							
33/76	If $A = \begin{pmatrix} a & b \\ c & d \end{pmatrix}$ and $B = \begin{pmatrix} d & -b \\ -c & a \end{pmatrix}$ , then B is a/an_matrix of A (0.06, 0.09, 0.16, 0.94, 0.91)							
(a) in	verse (b) transpose (c) ad joint (d) non-singular							
34/77	If two matrices are multiplied, the numbers of columns of the resultant matrix are equal							
	(0.38, 0.09, 0.30, 0.87, 0.85)							
	(a) rows of the first matrix (b) columns of the first matrix							
	(c) rows of the second matrix (d) columns of the second matrix							
	Geometry							
35/78	The pioneers of Geometry are (-0.87, 0.10, 0.20, 0.84, 1.03)							
	(a) Indians (b) Chinese (c) Egyptians (d) Germans							
36/79	Euclid wrote an important book on Geometry entitled (-0.70, 0.10, 0.25, 1.04, 0.92)							
(a) C	Concepts. (b) Elements. (c) Fundamentals. (d) Theorems.							
37/80	If an assumption is correct, then the results obtained from it are also correct. This principle is							
	governed by method.							
	(0.24, 0.09, 0.32, 0.79, 0.86)							
(a) ana	lysis (b) synthesis (c) analytic-synthetic (d) reductio-Ad-Absurdum							
38/81	The number of elements for proving a geometrical theorem is $(0.42, 0.09, 0.23, 0.86, 1.06)$							
	(a) 4 (b) 5 (c) 6 (d) 7							
39/82	40° and 50° are angles of each other (-0.48, 0.10, 0.12, 0.94, 0.87)							
	(a) supplementary (b) complementary (c) vertical (d) corresponding							
40/83	In any triangle, at least two angles are angles. (0.24, 0.09, 0.20, 0.88, 1.03)							
	(a) acute (b) right (c) obtuse (d) interior							
41/84	The medians of a triangle intersect each other in a ratio of (0.33, 0.09, 0.15, 0.97, 0.76)							
	(a) 1:4 (b) 1:3 (c) 1:2 (d) 1:1							
42/85	Each angle of an equilateral triangle is in measure. (-0.70, 0.10, 0.18, 0.94, 0.91)							
	(a) $30^{\circ}$ (b) $45^{\circ}$ (c) $60^{\circ}$ (d) $75^{\circ}$							
43/86	In a/an triangle, all the three right bisectors of the sides are concurrent at a point lying							
	within the triangle.							
	(0.51, 0.09, 0.30, 1.05, 0.08)							
(a) 1	right angled (b) acute angled (c) obtuse angled (d) all the three							